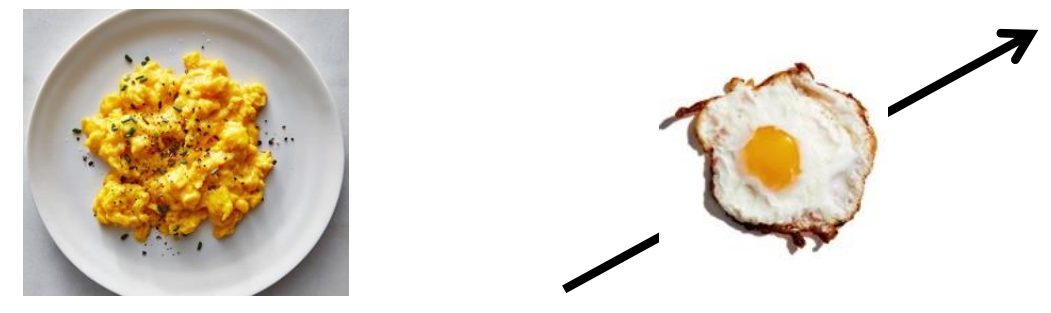


Spatial Extremal Dependence in Natural Hazard Footprints

Laura C. Dawkins¹ & David B. Stephenson²

1. Motivation

- Re-insurers use **catastrophe models** to estimate risk of natural hazards for insurance pricing. These models require **synthetic simulations** of natural hazard events for validation
- We develop a **statistical model** to represent the relevant properties of the hazard footprint, allowing for **very quick simulation** of synthetic events
- Hazard losses are sensitive to the **dependency between extreme values of the hazard variable at different spatial locations**. The statistical model must therefore accurately represent this **extremal dependence**



Modelling the scrambled egg (hazard footprint), rather than the fried egg (weather system)!

2. Application

Windstorm Footprint:

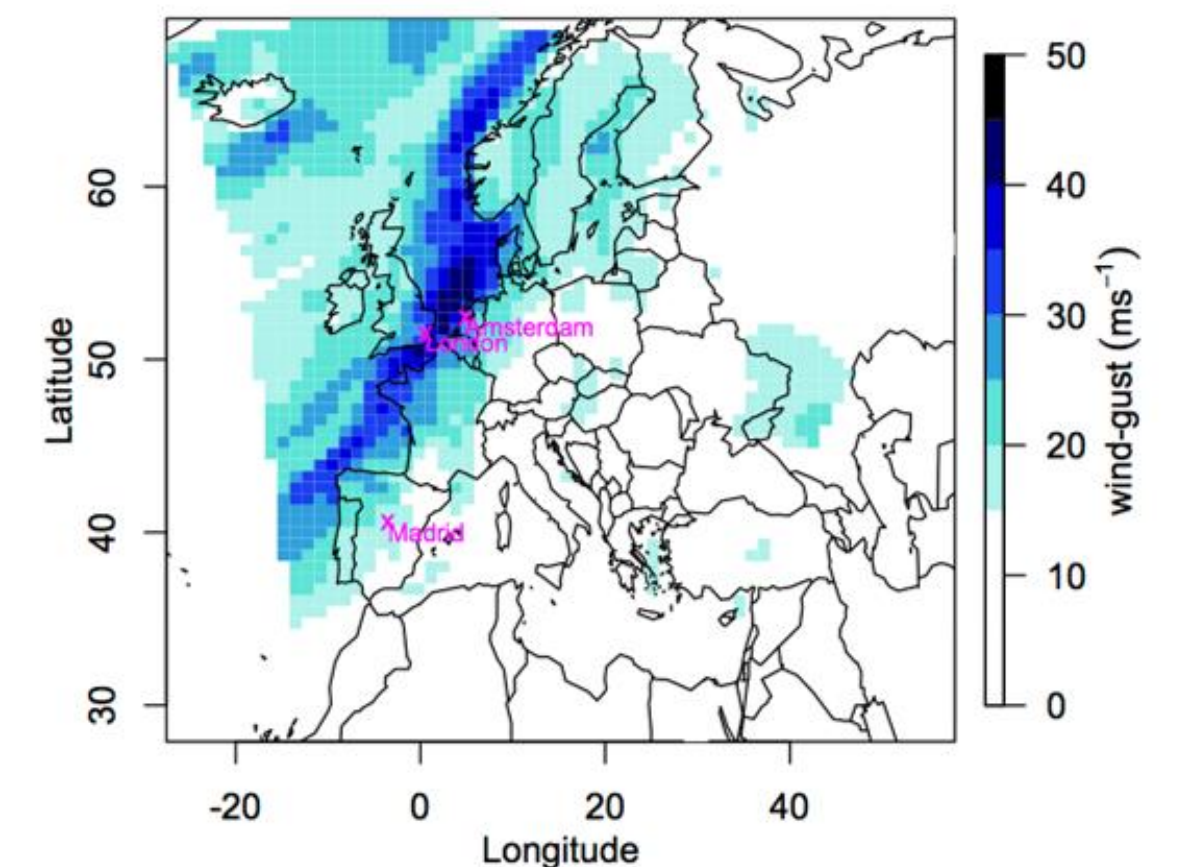
The maximum 3-second wind-gust speed at each location in the 72 hour period covering the passage of the storm, centred on the time at which the maximum wind speed over land occurs

Data set:

- 6103 storms, from extended winters (October-March) 1979-2014
- Dynamical downscaling of ERA-Interim reanalysis data using the Met Office 25km resolution North Atlantic-European operational numerical weather prediction model (Roberts et al., 2014)

Footprint for the Great Storm of October '87 (15th – 17th October 1987)

Caused £4 billion of insured losses in Europe

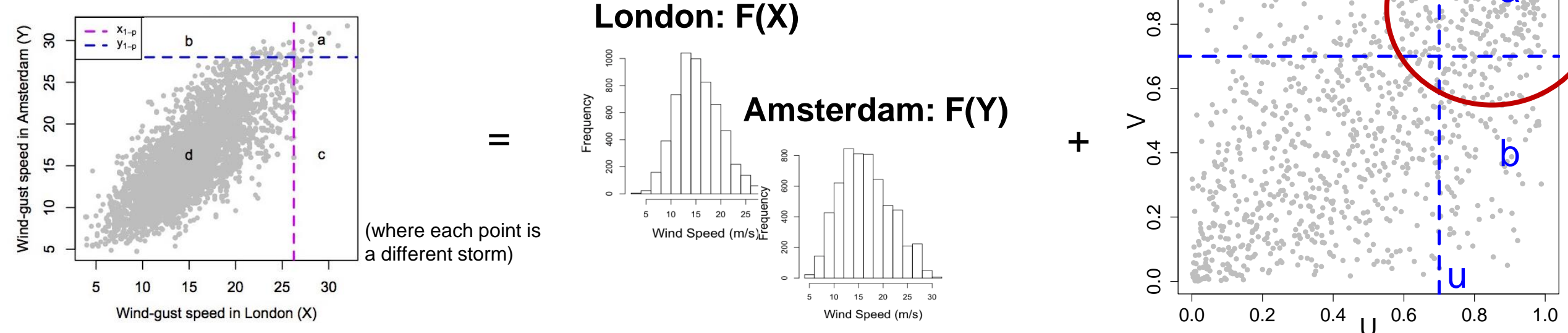


3. Exploring Extremal Dependence

The **joint distribution** of footprint wind-gust speeds at two locations, denoted X and Y, can be thought of in terms of...

...the distributions of gusts at each location ...

... and their mutual dependence



Either **Asymptotic Independence**: Largest values of the variable rarely occur together

Or **Asymptotic Dependence**: Large values of the variables tend to occur simultaneously

Method 1: Extremal Dependence Measures (Coles et al. 1999)

$$\chi(u) = \frac{\Pr(V > u | U > u)}{\Pr(V > u, U > u) / \Pr(U > u)} = \frac{a}{a+b}$$

$$\bar{\chi}(u) = \frac{\log((a+b)/n)}{\log(a/n)} - 1$$

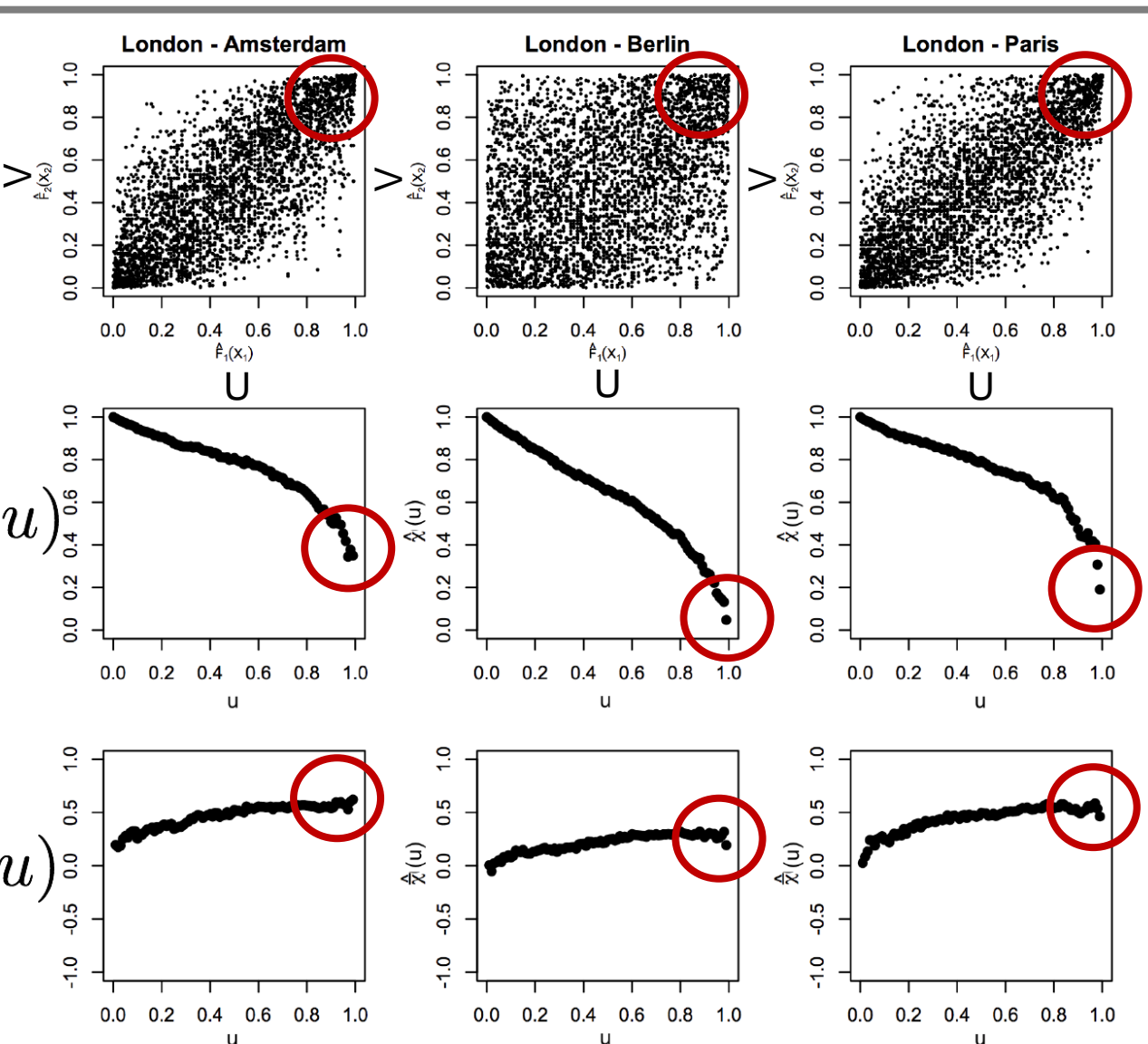
Empirical estimates suggest **asymptotic independence**, but these cannot represent the asymptotic limit $u \rightarrow 1$

Asymptotic Independence:

$$\lim_{u \rightarrow 1} \chi(u) = 0 \quad \lim_{u \rightarrow 1} \bar{\chi}(u) < 1$$

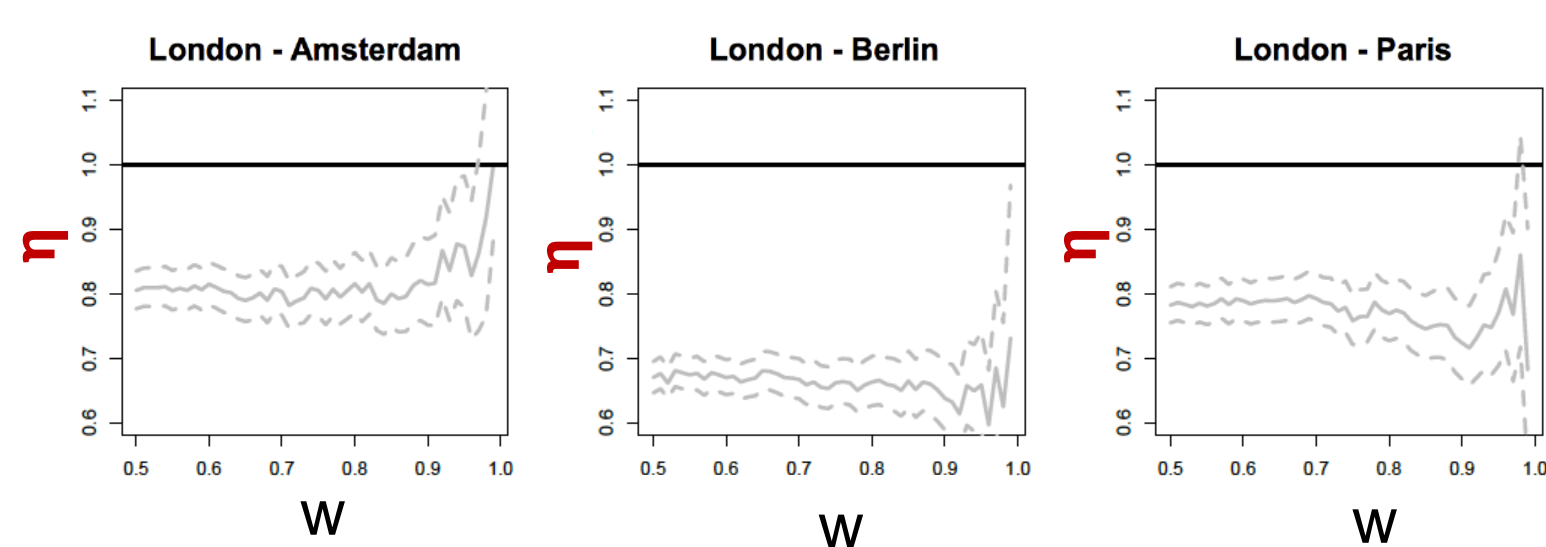
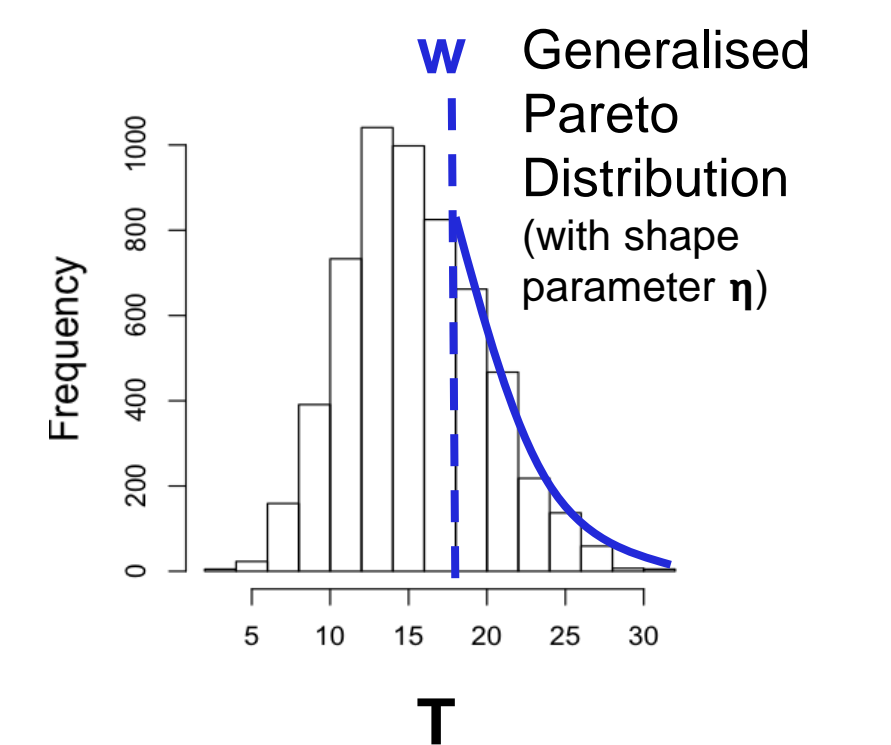
Asymptotic Dependence:

$$\lim_{u \rightarrow 1} \chi(u) > 0 \quad \lim_{u \rightarrow 1} \bar{\chi}(u) = 1$$



Method 2: Coefficient of Tail Dependence (Ledford and Tawn, 1996)

- Transform wind gust speeds at each location to follow a Laplace distribution;
- For each pair of observations take the minimum, denoted T;
- Fit a univariate extreme value Generalised Pareto Distribution to T, above a high threshold w;
- This distribution has three parameters: location, scale and shape;
- The shape parameter, η , of this distribution characterises extremal dependence, representative of the limit $u \rightarrow 1$;
- Repeat for a range of values of w to explore sensitivity to this choice of threshold



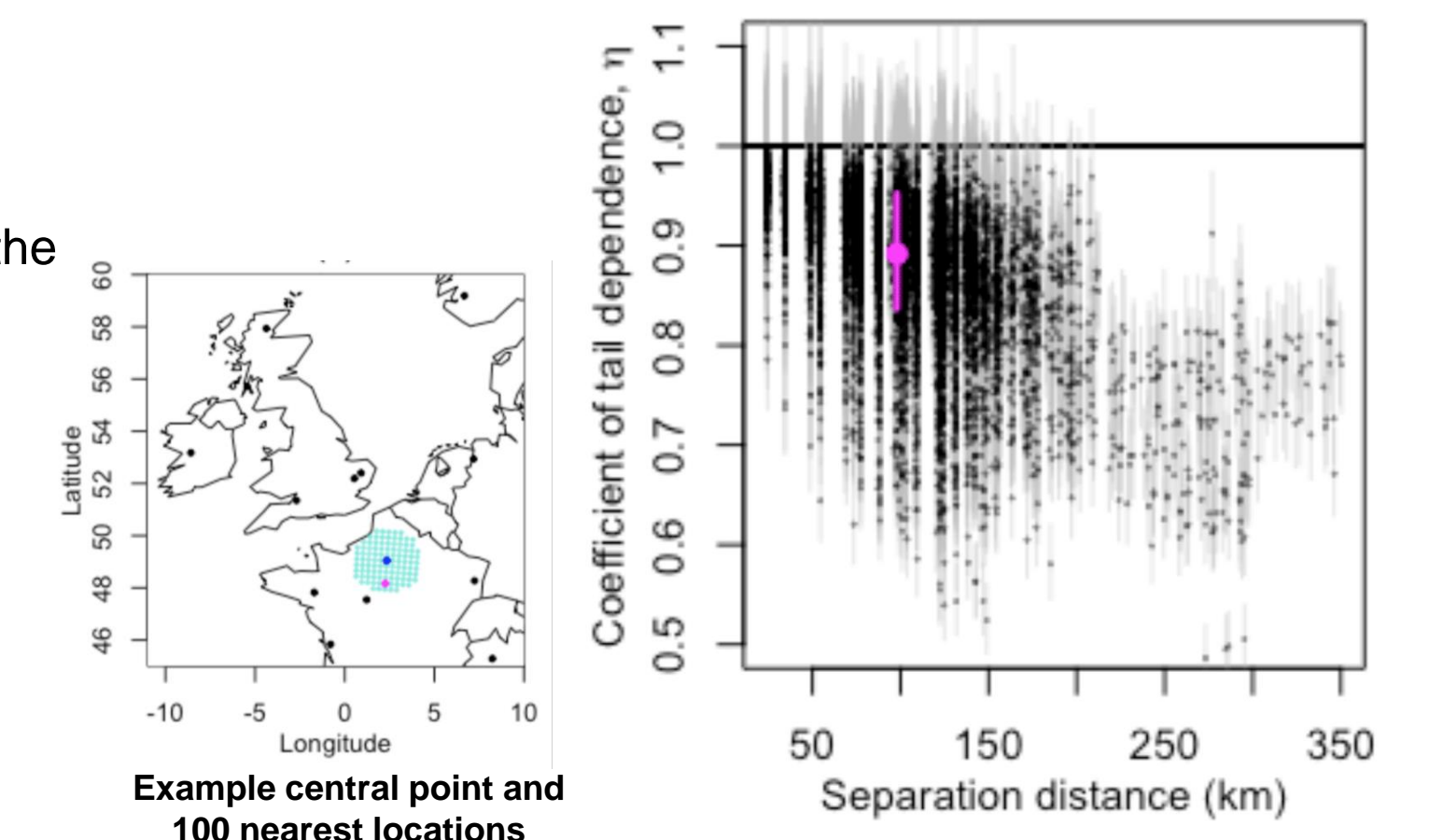
$\eta < 1$
Asymptotic independence

$\eta = 1$
Asymptotic dependence

For these 3 pairs of locations $\eta < 1$ for most w's in the range 0.5-1, suggesting **asymptotic independence**. What about other pairs of locations?

- Take a stratified sample of 100 locations in the European footprint domain
- Estimate η for each of these locations, paired with the 100 nearest locations
- Explore η for increasing separation distance

- Indication of **extremal dependence** for some very nearby locations
- Asymptotic independence is the dominant extremal dependence structure** across the spatial domain (200 – 3500km all $\eta < 1$)



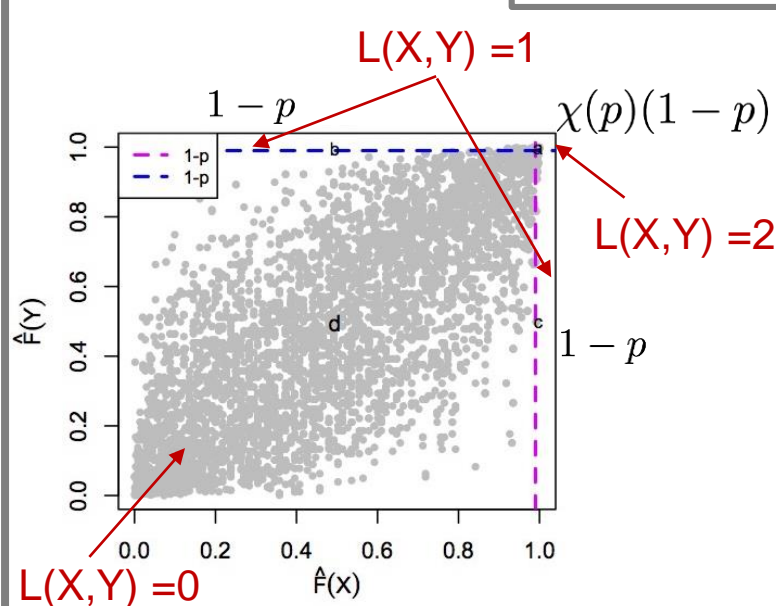
4. Natural Hazard Loss Diagnostic

Bivariate windstorm loss model (Roberts et al. 2014)

- Loss occurs when gusts exceed a **location specific loss threshold**, the **p=0.99 quantile** footprint wind gust speed (1979-2014)

$$L(X, Y) = 1(X > u(X)) + 1(Y > u(Y))$$

where 1(x) is an indicator function



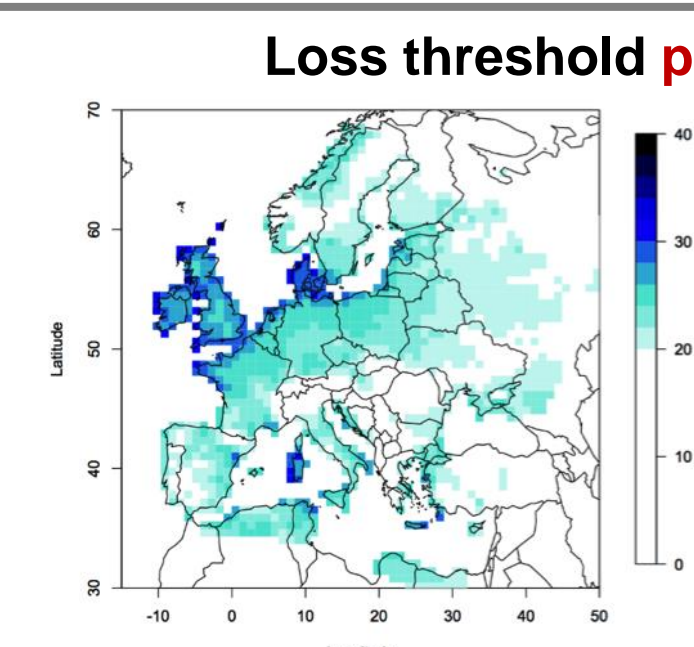
$$\chi(p) = \Pr(\hat{F}(Y) > p | \hat{F}(X) > p)$$

$$\Pr(L(X, Y) = 2) = \chi(p)(1-p)$$

$$\Pr(L(X, Y) = 1) = 2(1-\chi(p))(1-p)$$

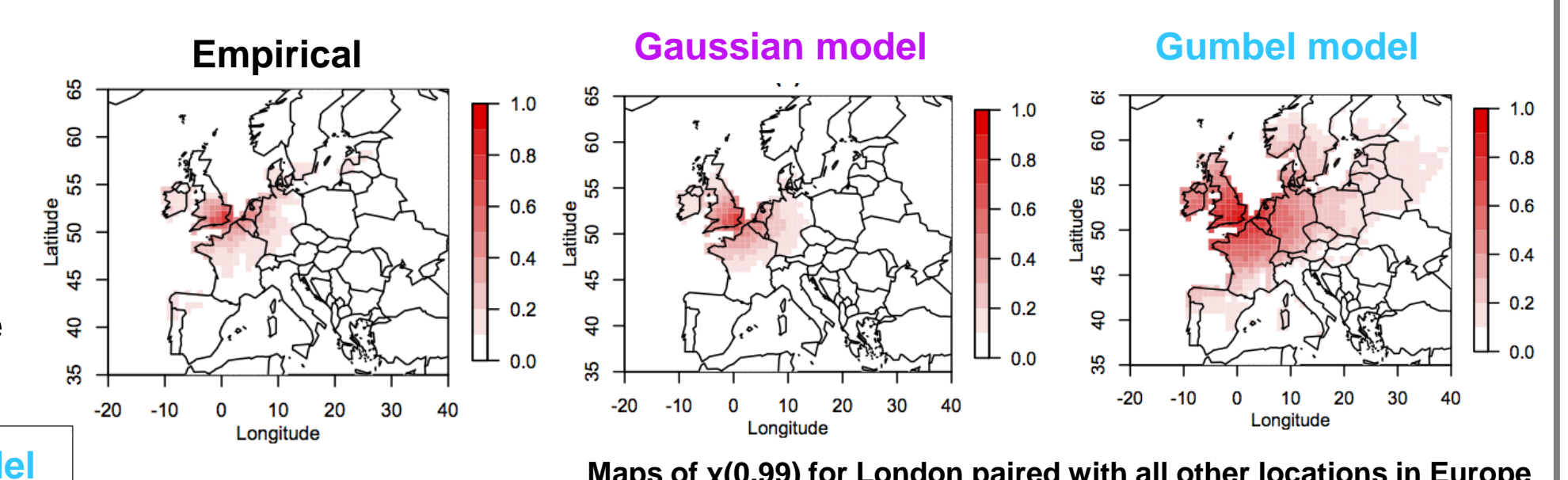
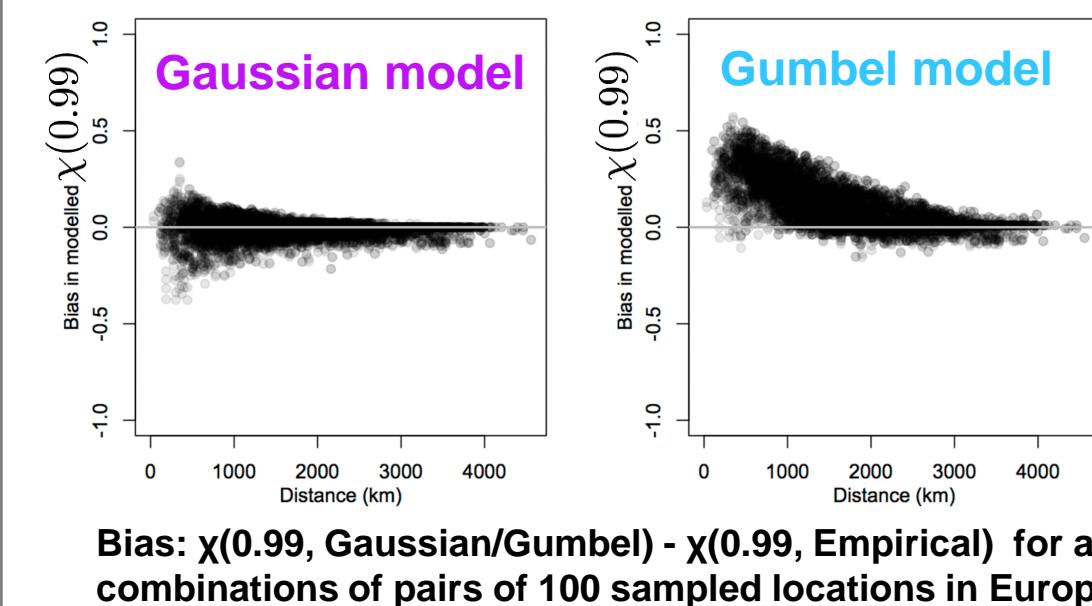
$$\Pr(L(X, Y) = 0) = (\chi(p)-1)(1-2p)$$

The windstorm **loss distribution** for a pair of locations can be expressed in terms of $\chi(p)$, where p=0.99 quantile wind gust at each location



Extremal dependence and Loss

- The **Gaussian** dependence model characterises asymptotic independence
- The **Gumbel** dependence model characterises asymptotic dependence



- The **Gaussian model (asymptotic independence)** represents the empirical joint losses well
- The Gumbel model (asymptotic dependence) over-estimates dependence between pairs of locations, particularly those <1000km apart, resulting in over-estimation of joint losses
- This gives further evidence of **asymptotic independence** in windstorm footprint wind gusts

5. Why are windstorm footprints asymptotically independent?

- Turbulent wind fields** can be efficiently and realistically simulated using **stochastic processes**
- This approach is widely used for many applications such as testing loads on new aircraft designs

- Such models assume that the Cartesian velocity components at location j: u_j and v_j such that $X_j = \sqrt{u_j^2 + v_j^2}$, are **independent Gaussian processes**, each with a prescribed spatial covariance function

- Since the individual velocity components are bivariate normal, they are **asymptotically independent** at different locations
- e.g. u_1 and u_2 are asymptotically independent when locations $j = 1$ and $j = 2$ differ, and likewise for v_1 and v_2

- Furthermore, it can be shown that the square of each velocity component is also **asymptotically independent**
- See Appendix of Dawkins and Stephenson (2018)

- The squared wind speeds at pairs of locations are sums of two such independent components: $(X_1^2, X_2^2) = (u_1^2 + v_1^2, u_2^2 + v_2^2)$
- Asymptotic independence** suggested by simulation in Dawkins and Stephenson (2018), now proven.

6. Conclusion

- Asymptotic independence** is the dominant form of extremal dependence in the 6103 dynamically-downscaled European windstorm footprints
- This is supported by the properties of stochastic models for turbulence
- The **Gaussian** model appears to fit the data well, which allows for **fast geostatistical simulation** of footprints
- For further details see Dawkins and Stephenson (2018)

References

Dawkins L. C. and Stephenson D. B. (2018). Quantification of extremal dependence in spatial natural hazard footprints: independence of windstorm gust speeds and its impact on aggregate losses. *Natural Hazards and Earth System Science*, 18, 2933-2949

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